

1 **Spatial Structure among the Geometric Earthworks of western Amazonia (Acre, Brazil)**

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12 **Highlights**

13 • Inferring socio-political organisation among Amazonian geometric earthwork builders is a
14 significant challenge.

15 • Point process modelling is adapted for detecting multi-scalar spatial structure in a sample of
16 geometric earthworks.

17 • Controlling for first- and second-order spatial dependency highlights likely critical scale of
18 spatial structure.

19 • This information provides the first robust regional estimates of possible territorial integration.

20 • Hierarchical (vertical) organisation is counter-indicated, suggesting geometric earthwork
21 building societies more likely had non-centralised or flat structures.

22

23 **Abstract**

24 Fieldwork and desk-based research in the western Amazon basin has led to an explosive growth
25 in the state of knowledge surrounding the pre-Columbian archaeology of this region.
26 Previously thought to be a sparsely occupied environment, archaeologists have recorded
27 hundreds of geometric earthworks between the Purús and Acre rivers in recent years, spurring
28 renewed interest in understanding the distribution, age, and function of these structures. A
29 challenge has been to identify possible relationships between sites and to place them in their
30 broader landscape setting. The precise spatial scale, relative importance of different factors,
31 and strength of any relationships that contributed to shaping their distributions remain an open
32 question. This paper develops and applies an explicitly spatial framework to address this
33 problem, drawing on a rich body of recent research in Acre state (Brazil) and advanced point

34 process modelling. The analytical approach, which is fully documented and reproducible with
35 the accompanying code, infers the factors affecting [geometric](#) earthwork distribution at
36 multiple spatial scales. This enables the first robust predictions of territorial integration in the
37 region, which is discussed context of extant archaeological models. The findings support the
38 interpretation that non-stratified societies [likely](#) occupied Acre during the late pre-Columbian
39 period.

40

41 **Keywords**

42 Amazonia, spatial statistics, territoriality, social complexity, South America, earthworks

43

44 **1. Introduction**

45 The geometric earthworks of the western Amazon basin have received substantial attention in
46 the pursuit of understanding their distribution, context, purpose, and meaning. First
47 documented in the early 1990s, and possibly earlier (Ranzi 2003; Ranzi et al. 2007; Dias 2006;
48 Schaan 2012), the known structures are located primarily in the interfluvial uplands of the
49 upper Purús and Acre rivers in the modern Brazilian state of Acre. Viewed in the context of
50 key debates in Amazonian archaeology that were coming to a head by the end of the 1990s and
51 early 2000s (Meggers 1992; Denevan 1996; Heckenberger et al. 1999; 2001; Meggers 2001;
52 Stahl 2002), geometric earthworks have attracted widespread interest as proof that indigenous
53 societies could thrive far from more agriculturally-attractive sedimentary floodplains (Schaan
54 2012; Watling et al. 2018). With time, well over 400 individual [geometric](#) earthworks, also
55 termed ditched enclosures or “geoglyphs”, have been documented through aerial and remote
56 sensing surveys (Ranzi et al. 2007; Schaan 2012; Saunaluoma 2012). This number increases
57 with every publication, in the course of which archaeologists have continued to add nuance as
58 new information is revealed. Notably, targeted efforts to understand [geometric](#) earthwork
59 chronology, function, and ecological context through excavation have yielded significant
60 insights into the societies that built them, as well as their land use practices (Saunaluoma and
61 Virtanen 2015; Watling et al. 2017). Archaeological ceramics recovered from [geometric](#)
62 earthworks suggest their embeddedness within a broader pan-Amazonian cultural milieu and a
63 probable ceremonial purpose in prestige-enhancing displays (Pärssinen et al. 2009; Schaan et
64 al. 2007; Schaan et al. 2010; Saunaluoma 2012). On the other hand, the relative scarcity of
65 cultural deposits argues against persistent occupation within [these sites](#). They are,
66 consequently, interpreted as unlikely to have functioned as settlements. Taken as a whole,
67 radiocarbon assays on geometric earthworks have yielded dates spanning at least two millennia

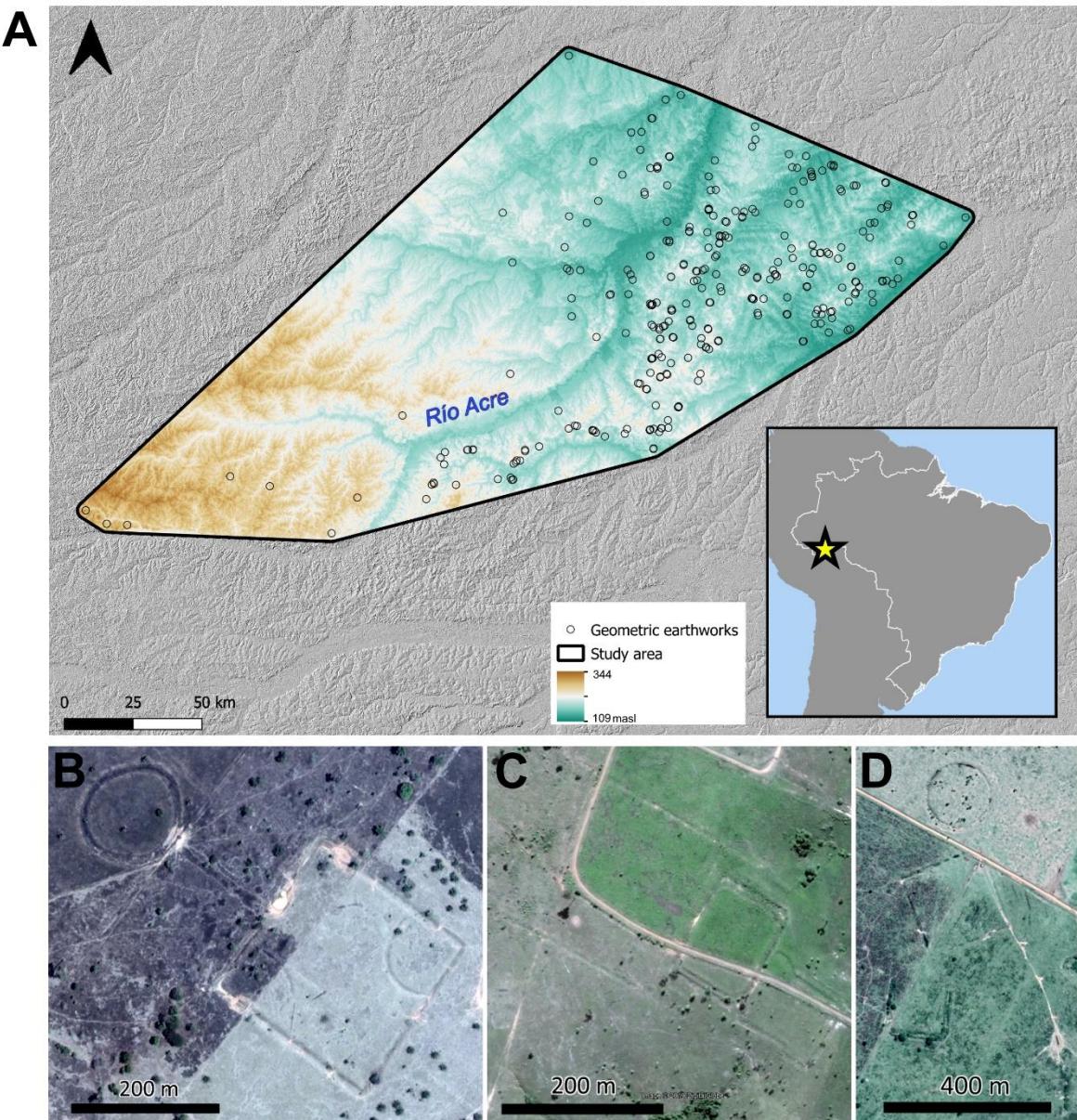


Fig. 1: Overview of geometric earthworks in the study area and location in Brazil. A: The study area comprises the convex hull of all earthworks located in Acre state. Note the spatial inhomogeneity of the point pattern in an east-to-west density cline. B: The Fazenda Atlântica site, showing internal subdivisions and inter-earthwork connections. C: Unnamed series of quadrangular earthworks in northern Acre, with multiple ditches and berms. D: The Fazenda Colorado site, showing spatial associations and alignments between circular and quadrangular earthworks. Source: ESRI & Google Earth Imagery.

from 1215-1009 cal BC to 1218-1291 cal AD (Ua-37238 and Hela-616, respectively, calibrated at 2σ with the SHCal13 curve, Saunaluoma and Schaan 2012; Hogg et al. 2013), although there are significant gaps across the set, as well as occasional unresolved inversions in individual soil profile ages (Watling et al. 2017). Terrestrial palaeoecological records reveal an apparent absence of large-scale forest clearance in the landscapes around geometric earthworks throughout their entire period of use (Watling et al. 2017). Nonetheless, microbotanical data from archaeological contexts provide evidence for subsistence practices that employed both

75 domesticated and wild food resources, in particular palm fruit and brazil nut, suggesting that a
76 mixed system of managed agroforestry with cultigens existed in the past (Watling et al. 2015;
77 Watling et al. 2018). Although circular and rectangular plans predominate among geometric
78 earthworks, a variety of shapes have been reported, including polygons, ovals, and D- and U-
79 shaped earthworks. Rarely, causeways or “avenues” connect earthworks, while the enclosures
80 themselves can have single or multiple ditches and berms. Within this relatively restricted
81 range of attributes, geometric earthworks nonetheless exhibit a high degree of formal
82 variability in their architectural characteristics and configurations. In certain rare cases,
83 earthworks were constructed either abutting or overlaying previous earthworks, possibly
84 indicating the long-term use of particular places and parts of the landscape. Schaan (2012), in
85 plotting the raw distribution of earthwork shapes and sizes, has suggested that these variables
86 might encode cultural preferences or chronological trends.

87

88 Major advances in knowledge on the geometric earthworks have occurred in a relatively short
89 span of time, speaking to a vigorous programmatic interest among scholars of Amazonian
90 indigenous history. Extensive research across the southern rim of the basin has attested to other,
91 occasionally contemporaneous, traditions of earthwork construction that performed a range of
92 domestic, ceremonial, and subsistence roles for the societies that erected them (Denevan 1963;
93 Arnold and Prettol 1988; Heckenberger 2005; Walker 2008; Erickson 2006; 2010; Lombardo
94 and Prümers 2010; Carson et al. 2014; Prümers 2014; Prümers and Jaimes Betancourt 2014;
95 De Souza et al. 2018). Archaeologists have also written extensively on the degree of social
96 complexity among the populations involved in the construction of the geometric earthworks in
97 the modern states of Acre, Rondônia, and Amazonas (Pärssinen et al. 2009; Schaan 2012;
98 Saunaluoma and Schaan 2012; Saunaluoma 2012; Saunaluoma and Virtanen 2015;
99 Saunaluoma et al. 2018; Watling et al. 2018). Taking a broad view of the above publications,
100 certain commonalities emerge and can be summarised here briefly. First, sizeable populations,
101 possibly totalling in the tens of thousands, are estimated to be responsible for the geometric
102 earthworks. Concordantly, some degree of supra-local organisation may well be behind
103 coordinating the number of people at, or close to, such an order of magnitude. The available
104 archaeological data does not indicate the widespread or ubiquitous settlement nucleation,
105 however, suggesting that settlement was dispersed over a relatively large area, or has at least
106 left a less obtrusive archaeological signal (Saunaluoma et al. 2018). Political authority, as such,
107 was likely to have been transient and derived from skilfully leveraging symbolic or ritual
108 capital in the context of ceremonial performance associated to geometric earthworks.

109

110 Analyses of the broader landscape of geometric earthworks have proposed social processes
111 behind their formation with direct reference to a model of organisational structure that
112 underscores the potential interplay between hierarchy and heterarchy (Crumley 1979; 1995).
113 The latter term implies power relations that are distributed, unranked, or marked in varying
114 ways depending on the scale or frame of reference, emphasising the relational nature of
115 authority derived through corporate mechanisms (Ashmore 2002; Morehart et al. 2018: 14;
116 Furholt et al. 2019). The term is not construed as inimical or opposed to the existence of
117 hierarchical, vertical, or ranked structures, which may co-exist or be negotiated alongside such
118 lateral relationships (DeMarrais 2013). Indeed, the existence of multiple equivalently valued
119 agencies or agendas in heterarchical configurations, and the societal tensions stemming from
120 the resulting dynamics, is suggested as a major driver of social differentiation among humans
121 (Furholt et al. 2019: 7). Prehistoric contexts across the globe provide examples for this pattern,
122 and there are compelling empirical reasons to entertain the possibility of their existence in
123 Amazonia too (Roosevelt 1999; Schaan 2012; Balée et al. 2014). Specifically, pre-Columbian
124 communities are envisioned as nominally autonomous yet possessing internal social rankings
125 that were occasionally brought to the fore. Materialised in Acre in the form of geometric
126 earthworks, the interplay between distinctive types of organisation in the pre-Columbian period
127 is cited as the dynamic that drove the formation of “complex” social formations (Schaan et al.
128 2010; Schaan 2012; Saunaluoma et al. 2018).

129

130 *Formalising spatial relationships*

131 Interpretative models of landscapes and land use provide a useful point of departure for
132 developing the present spatial analytical approach toward geometric earthworks. This paper
133 contends that directly investigating spatial structure at multiple scales may help elucidate the
134 dynamics responsible for the formation of the archaeological landscape of Acre (Crumley
135 1979: 144). In particular, I draw on the suggestion that regional heterarchies are an inherently
136 multiscalar, spatial phenomena (Crumley 1979, 1995; Ashmore 2002). Although formal tests
137 of this assertion have been slow to materialise (Ashmore 2002: 1177; Smith and Schreiber
138 2005; Carballo et al. 2014), modern point process modelling offers the opportunity to test for
139 the presence and type of spatial behaviour within a large sample of significant monumental
140 sites (Baddeley and Turner 2005; Baddeley et al. 2016). The vast area and chronological span
141 over which geometric earthworks occur pose a challenge to generalising as to their precise role
142 among the pre-Columbian societies that constructed them. With the data currently available,

143 we cannot at present preclude that both function and meaning also changed substantially over
144 the two millennia or more of their existence. **Multiscalar** point pattern analysis has emerged as
145 one of the solutions for direct analyses of the complexity of the archaeological record (Bevan
146 et al. 2013; Eve and Crema 2014; Riris 2017; Dreslerová and Demján 2019). These methods
147 offer an advantage over fixed-scale descriptive spatial statistics, such as nearest neighbour,
148 which are easily confounded by the spatially variable nature of much archaeological data (Riris
149 2017; Carrero-Pazos et al. 2019). In another sense, computational modelling also offers the
150 opportunity to test the myriad interpretations of the archaeological record. Here, incorporating
151 a range of spatial effects and processes into formal models, including what are usually termed
152 “first-order” (exogenously induced variables influencing the density of points) and “second-
153 order” (the density of points as affected by the endogenous configurations of other points)
154 effects, put these types of analyses on a robust footing for inferring socio-political dynamics.
155 In particular, what was the spatial scale of integration among geometric earthwork building
156 societies? To what extent does spatial structure index social structure? Is it possible to settle on
157 just one scale, singular, or do patterns of import occur at multiple spatial scales? These
158 questions reflect notable gaps in knowledge on the presence and strength of spatial
159 relationships between geometric earthworks, which this paper will attempt to address
160 employing a formal modelling approach.

161
162 Archaeologists observe marked consistencies among the settings of geometric earthworks
163 (Schaan et al. 2010; Schaan 2012; Saunaluoma et al. 2018). Even a cursory look at the overall
164 distribution indicates that geometric earthworks are far from randomly distributed (**Fig. 1**). To
165 briefly summarise the locational characteristics that appear repeatedly and consistently across
166 the abovementioned literature: geometric earthworks are typically located at elevation within
167 the overall gently rolling topography of the Purús and Acre interfluves. Moreover, their
168 placement also tends to be on prominent bluffs relative to their surroundings, as opposed to
169 simply at altitude in absolute terms. Finally, while water availability is not typically thought to
170 be a factor despite the seasonally dry (monsoonal) climate of the region, proximity to large,
171 perennial watercourses, as potential arteries of communication and transport, likely exert some
172 influence on the placement of these sites. Next, and temporarily leaving the landscape to one
173 side, a closer look at the locations of geometric earthworks alone may indicate that they form
174 clusters separated by relatively short (~5 km or less) or alternatively at much longer ranges
175 (>10 km) too. In other words, monument locations may also depend in some way on the
176 locations of others. Notably, Schaan (2012: 154-155) explicitly connects site size to labour

177 mobilisation and suggests the existence of territorial interaction between clusters at a regional
178 scale, centred on the locations of the largest of geometric earthworks. In this sense, and with
179 reference to the inferred presence of complex heterarchical organisation here, spatial
180 distributions have long been interpreted as indexing social relations.

181

182 A principal strength of the method developed and applied here is the explicit recognition that
183 the factors that drive spatial phenomena (here, monument location) can also vary spatially. A
184 basic inspection of geometric earthworks distributed within Acre indicates that, like many
185 empirical point patterns, spatial inhomogeneity cannot be ruled out (Fig. 1). Previous locational
186 models employing similar data have largely employed methods that assume a homogeneous
187 point pattern intensity (McMichael et al. 2014a; 2014b; De Souza et al. 2018). This paper
188 attempts to remedy this by examining the variability in spatial patterning (inhomogeneous
189 dispersion and/or aggregation) at multiple scales simultaneously, while controlling for both
190 induced effects and between-point interaction. This advance in method bridges the correlative
191 (first-order) and relational (second-order) spatial properties of the archaeological record, which
192 are frequently key factors in data that is fundamentally human in nature.

193

194 In summary, this paper uses point process modelling to explicitly control for the effects of
195 external variables on the distribution of geometric earthworks, which are drawn from the
196 literature. The approach also introduces a second-order (inter-site) interaction term, which
197 together with the initial model successfully describes the majority of spatial structure among
198 the geometric earthworks of Acre. Together, this multivariate model evaluates the relative
199 influence of concurrent spatial phenomena on the distribution of geometric earthworks in a
200 robust statistical framework. Additional tests on the fitted models reveal a statistically
201 significant clustering threshold, identifying empirically-supported groups of sites in the
202 process. A Monte Carlo test of the distribution of geometric earthwork sizes within and
203 between these groups (Carrero-Pazos et al. 2019) is adapted to investigate if intra-cluster site
204 hierarchies are present. Summarising, based on present information, the model-testing
205 approach indicates that the populations responsible for the creation of geometric earthworks in
206 western Amazonia are highly unlikely to have been hierarchically organised either at a regional
207 scale, or within groups of sites.

208

209 **2. Materials and Methods**

210 This study focuses primarily on pre-Columbian geometric earthworks in Acre state, Brazil, the
211 core area reported in the literature (Schaan 2012; Saunaluoma and Virtanen 2015; Watling et
212 al. 2018). **Hereafter, all references to earthworks refer to geometric earthworks located in Acre**
213 **unless otherwise noted.** The reported locations of archaeological sites were gathered through a
214 literature survey, which formed the initial dataset. All earthwork locations were also cross-
215 referenced with public domain satellite imagery available through the ESRI ArcGIS Online
216 platform and Google Earth. Following analytical and interpretative precedent by Schaan
217 (2012), earthwork surface area was directly estimated from the available imagery in the course
218 of data collection. **Surface area was consistently measured from the outer perimeter of**
219 **earthwork banks or berms and the area enclosed by ditches therefore contributes to the total.**
220 In rare cases where multiple enclosures are connected through ‘avenues’, and sequences of
221 construction cannot be discerned, their combined area is reported. Furthermore, for the handful
222 of earthworks (<5) that *directly* abut or overlay each other, and where the sequence of
223 construction cannot be discerned from remotely sensed imagery alone, their location is reported
224 as a single centroid of the combined outlines. Other aspects of geometric earthwork
225 monumentality, such as berm height and ditch depth or indices of architectural complexity,
226 may conceivably be relevant to understanding the energetics and relative configuration of the
227 archaeological landscape of Acre. **In this paper, I limit analysis to area enclosed as a first order**
228 **approximation of earthwork size as a possible index of elaboration (Schaan 2012).**

229

230 Data collection resulted in a total of 420 point locations that could be derived directly from
231 published maps, with associated information. The majority of **earthworks** ($n = 312$) lie within
232 the state of Acre, where agribusiness-led deforestation, centred on the town of Rio Branco, is
233 most extensive. The remaining 108 earthworks were encountered in Rondônia and Amazonas
234 states. Absent widely available canopy-penetrating remote sensing technology (Prümers and
235 Jaimes Betancourt 2014), it must be cautioned that the representativity of the sample is
236 essentially dictated by the extent of deforestation. As this process is decidedly more advanced
237 in Acre compared to neighbouring states, the data affords improved control and explanatory
238 power over the areas of Amazonas and Rondônia where forest removal is less complete. To
239 avoid spurious results due to uneven sampling, this study is restricted to the area formed by the
240 convex hull of earthworks within Acre, buffered by 3 km to mitigate edge effects. This
241 approach withholds sites located in Rondônia and Amazonas from the study, pending
242 verification of the extent to which this subset of western Amazonian earthworks is
243 representative of a broader pattern **beyond Acre.**

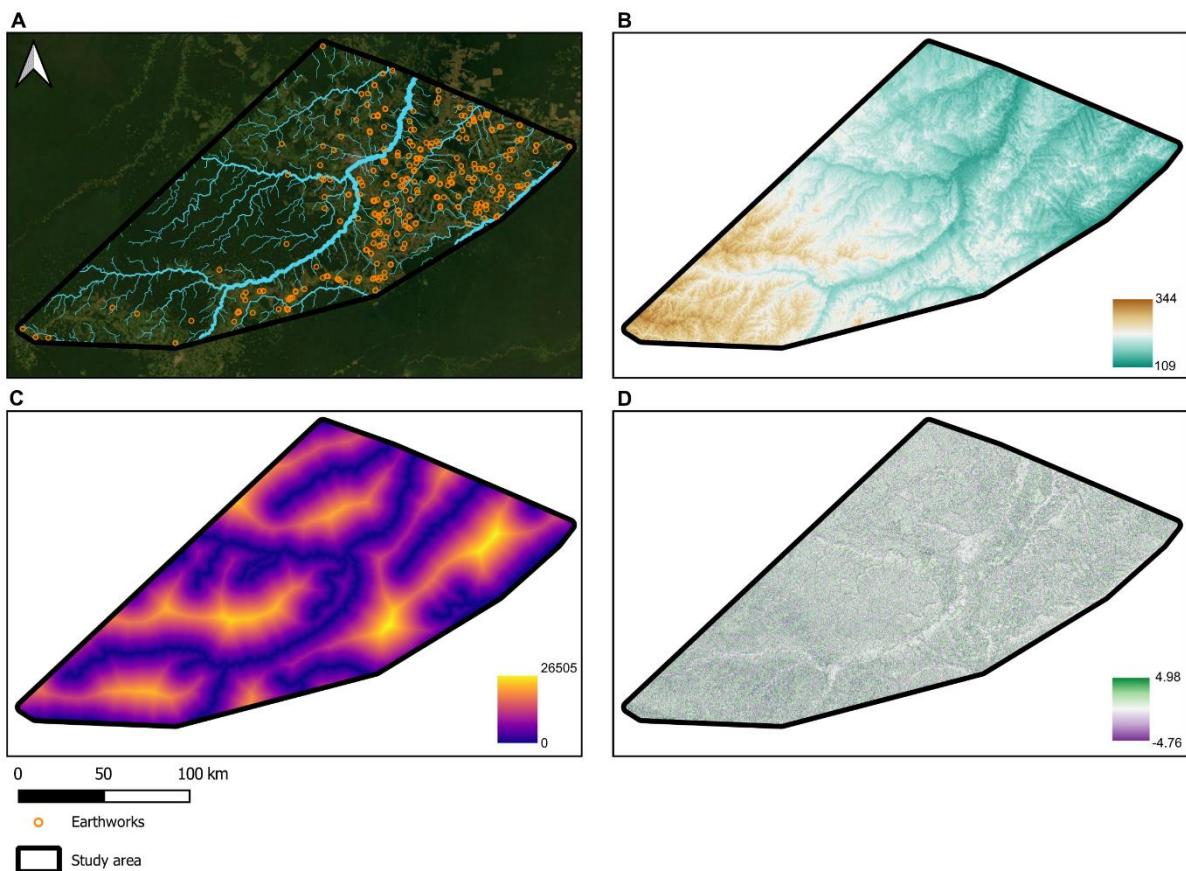


Fig. 2: Earthwork distribution and covariate data. A: Earthwork locations within the study area overlaid on satellite imagery, B: elevation raster (metres above sea level) derived from SRTM data, C: Euclidean distance (metres) from permanent watercourses, D: topographical position index (TPI, dimensionless).

As noted, point pattern analysis distinguishes primarily between first- and second-order spatial trends, that is, structure induced by external processes, and the degree of attraction or repulsion between points within a pattern (Diggle 2013). Bearing in mind the abovementioned explanatory frameworks for the distribution of Amazonian earthworks, this paper leverages modern spatial statistics and point process modelling to evaluate the degree to which archaeological patterns might be influenced by both induced *and* internal spatial structure concurrently. The principal findings are briefly summarised here, and the procedure is fully described in the accompanying code (**Supplementary Information**). Employing archaeological interpretations as a point of departure, the approach initially builds a model to control for the presence and strength of first-order spatial trends among Amazonian *geometric* earthworks. Next, a second-order interaction term is incorporated to account for spatial interaction not explained by the null and baseline first-order models. The results and insight gained from the modelling exercise detailed below reveals yet further statistically significant spatial structure that neither model fully accounts for, individually or in tandem. Based on these

259 findings, a combination of machine learning algorithms (Ester et al. 1996) and additional
260 spatial statistical analysis integrating earthwork attribute data (Carrero-Pazos et al. 2019),
261 yields robust estimates for possible western Amazonian territorial units. Moreover, this
262 procedure conservatively identifies zones where archaeological data is presently lacking
263 enough detail to make confident inferences, in particular where sites are exceptionally sparse.

264

265 *Fitting covariate spatial data*

266 The covariate spatial data (**Fig. 2**) are consistently highlighted in the archaeological literature
267 as relevant to understanding geometric earthwork distributions across Acre (Schaan et al. 2010;
268 Schaan 2012; Saunaluoma et al. 2018). Here, I assess the interplay and relative influence, if
269 any, of each dataset. First, a multicollinearity test with Pearson's R reveals that there are no
270 collinear effects between the datasets, indicating that they are suitable as independent model
271 terms. Next, the observed point pattern of geometric earthworks is fitted to a Poisson point
272 process model, equivalent to an initial null hypothesis of complete spatial randomness (CSR).
273 At the same, using the Bayesian Information Criterion (BIC) as a guideline for information
274 gain per added model term, a stepwise model selection procedure rejects none of the covariates
275 in isolation or in combination. The BIC suggests that a combination of all three covariates
276 provides the best starting point for modelling earthwork locations. Together, elevation,
277 topographical prominence, distance from water, and the assumed Poisson process comprise the
278 multivariate first-order model. The spatial intensity (density) of archaeological sites in Acre,
279 as described by the interaction of these parameters, is a function of their combined maximum
280 logistic likelihood. Third, to evaluate the possible effects of interaction at a landscape level
281 between earthworks, a second-order term is incorporated into the fitted first-order model to
282 generate a combined model. Such a model statistically accounts for induced and internal spatial
283 dynamics simultaneously. Here, following precedent for handling archaeological data in this
284 manner (Bevan et al. 2013), the area-interaction point process model is advanced to replace
285 CSR as the null and to capture the range of observed variability in the empirical point pattern.
286 Area-interaction is attractive due to its relative flexibility, generality, and lack of complex free
287 parameters. Alternatives such as the Hard Core, Matérn cluster, or Thomas process describe
288 processes that presume rather more about the formation of geometric earthwork distributions
289 than is presently known (Baddeley and Turner 2005). To obtain the best fit for the second-
290 order term, the joint probability distribution is estimated by log pseudolikelihood.

291

Finally, the goodness-of-fit of each of the three models is assessed comparatively employing the pair correlation function. This statistic has some especial advantages in the context of the goals of this study. For fitted models and point patterns, the statistic estimates the presence and degree of spatial clustering and repulsion at multiple scales within the window of observation. Unlike the classic Ripley's K statistic, which employs discs to count and measure spatial structure, the pair correlation function employs rings with outer and inner boundaries. Consequently, patterning at different scales is more easily discerned visually (Stoyan and Stoyan 1994; Baddeley et al. 2016). The fact that the function does not accumulate also provides a more intuitive result (Wiegand and Moloney 2004: 225; Jacquemyn et al: 2007, 451). Rather than the more common spatially uniform version of this statistic, the *inhomogeneous* variant used here accounts for differences in the density of the point pattern across the study. In other words, it assumes that the point pattern has an underlying spatial trend. Density-dependent variation is a well-known quality of earthwork locations (Schaan 2012; Saunaluoma et al. 2018) and controlling for this source of spatial structure can be considered an initial step in uncovering the possible drivers of the formation of the archaeological record.

To illustrate the critical scales where geometric earthworks attract or inhibit one another, significance bands are derived via Monte Carlo simulations of point patterns drawn from the null (random), first-, and second order models. Envelopes are generated by simulating 999 point patterns for each application of the pair-correlation function to the fitted models (Baddeley and Turner 2005; Jacquemyn et al. 2007: 451). Rather than expressing confidence to a chosen level (e.g. 95%), they estimate the likelihood of incorrectly rejecting the null hypothesis of no spatial structure as a function of the number of simulated runs. For the envelopes generated here, this value (α) corresponds to 0.002, which can be considered highly robust for the purposes of the modelling exercise. It must be acknowledged that the null model of complete spatial randomness (CSR) is unlikely to adequately reflect the empirical data, but it is nonetheless presented for completeness. As described above, the first-order model also assumes an underlying Poisson distribution, while the second-order model replaces this with an Area Interaction term. The workflow is fully implemented as R code in the supplementary information, with accompanying archaeological data, and draws extensively on functions in the package 'spatstat' version 1.60-1 (Baddeley and Turner 2005).

3. Spatial point process modelling

326 Spatial structure between geometric earthworks is well-illustrated by the initial application of
 327 the pair correlation function to the point pattern (**Fig. 3, top left**). Attending to the first
 328 maximum and minimum outside the simulation envelope, these values correspond to the typical
 329 range at which points are most strongly attracted or repelled by one another, respectively
 330 (Strand et al. 2007). For Acre earthworks, this occurs below 500 m for clustering and at
 331 approximately 13000 m for dispersion, with the observed value falling just outside the envelope

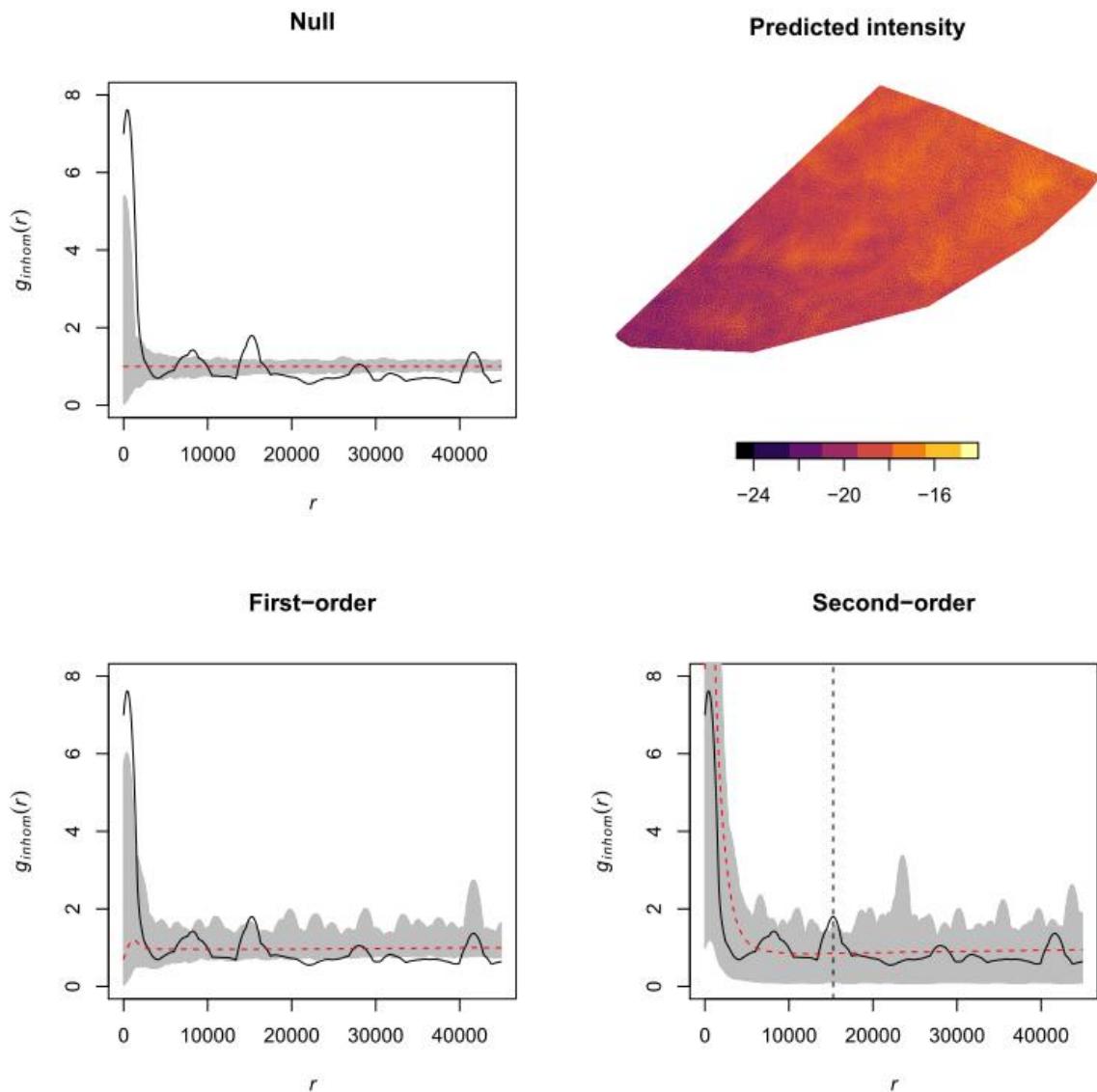


Fig. 3: Results of point process modelling. *Null*: The pair correlation function estimated on an assumption of complete spatial randomness, illustrating statistically significant clustering and repulsion that alternate at varying ranges (x-axes in metres). *Predicted intensity*: log-odds prediction surface of earthworks based on elevation, distance from major watercourses, and topographical prominence. *First-order*: the pair correlation function with a critical envelope conditioned on the covariate data as first-order variables, accounting for a greater range of spatial variability. *Second-order*: the final model, incorporating a point interaction term in the first-order model, accounting for all spatial variability except significant clustering at $r \approx 15000$ m (vertical line). Y-axes have been rescaled to a common range.

332 at this range. This initial finding suggests a variety of spatial processes may be at play, inducing
333 strong short-range clustering and somewhat weaker longer-range repulsion between
334 earthworks. Without further supporting information, such as fitted covariate data or point
335 process models beyond the null expectation of CSR, it must be underscored that this baseline
336 assessment lacks sufficient context to adequately explore the implications of these findings in
337 further detail. The analysis now turns to fitting the point pattern to covariate datasets derived
338 from close readings of the archaeological literature.

339

340 *First-order model*

341 The first-order model (**Fig. 3, bottom left**) provides a much improved fit over the CSR null
342 model, controlling for most of the weaker signals of spatial autocorrelation at all ranges from
343 >40 km to below 10 km. All three covariate datasets (**Table 1**) were found to be highly
344 significant. Nonetheless, the strong patterns of short-range clustering and long-range repulsion
345 of earthworks (beyond ~22 km) persist in this iteration. The predicted intensity (log-odds of
346 points per unit of area) from this model (**Fig. 3, top right**) provides an illustrative summary of
347 the degree to which this model explains spatial structure among geometric earthworks.
348 Although clearly spatially variable (inhomogeneous), the predicted spatial trend is weak, with
349 few zones showing values at the high end of the scale. Consequently, in agreement with the
350 results of the pair correlation function, the fitted first-order model indicates that a relatively
351 large proportion of spatial variation remains unaccounted for, in spite of the overall improved
352 fit. Rather than risk overfitting this model with covariate data beyond what is supported by
353 archaeological inference, the next step directly incorporates a model term to control for
354 between-point interaction, noting the exceptionally strong clustering found at ~450 m and
355 below.

356

357 **Table 1: Fitted covariate datasets for the first-order model.**

| Covariates | Estimate | Standard error | Z value | Significant |
|--------------------------|-------------|----------------|----------|-------------|
| (Intercept) | -15.6934100 | 0.408140 | -38.4511 | *** |
| Elevation | -0.0183163 | 0.002307 | -7.93976 | *** |
| Topographical prominence | 0.9620874 | 0.085621 | 11.23655 | *** |
| Distance to watercourses | 0.0000689 | 0.000011 | 6.341012 | *** |

358

359 *Second-order model*

360 The area-interaction process, as outlined above, is fitted by maximising the profile information
361 entropy with respect to a series of fixed spatial interaction distances (Baddeley and van
362 Lieshout 1995; Baddeley et al. 2016). The analysis is permuted in 100-metre intervals from
363 100 m to 2500 m to produce 25 candidate models, rather than impose a single arbitrary
364 interaction value. According to Akaike's Information Criterion, this procedure provides a best
365 fit for the area interaction parameter at $r = 2100$ meters. The second model also incorporates
366 the first-order model as a trend parameter. This combined first- and second-order model thus
367 successfully controls for both exogenous and endogenous processes within the point pattern of
368 geometric earthworks. Conditioning the simulation envelope for the pair correlation function
369 on the combined null model (**Fig. 3**, bottom right) illustrates the effect of incorporating
370 between-point interaction. The results reveal most variability in spatial structure among
371 geometric earthworks can be explained by controlling for these factors. While models with a
372 better goodness-of-fit could be derived from alternative readings of the archaeological record,
373 the data is nonetheless satisfactorily explained by the combined model. The model terms are
374 derived from close readings of the archaeological record, which reduces need to develop a more
375 complex model that risks overfitting against the data.

376

377 A notable exception of spatial structure for which the combined model does not account is the
378 presence of statistically significant clustering in the range $\sim 14500 - 15900$ at a significance
379 level of 0.002 (above), with a local maximum at $r = 15264$ m. This indicates spatial patterning
380 among geometric earthworks that may be particularly relevant for understanding distributional
381 trends of broader importance among the earthworks analysed here and, ultimately, the decision-
382 making processes that led to the formation of the archaeological landscape of Acre. This
383 information is now used to draw further inferences on landscape-level patterns among
384 geometric earthworks through further tests of their properties and configurations.

385

386 **4. Geometric earthwork clusters and size ranking**

387

388 Building a point process model of earthworks in Acre has proven informative in several
389 regards. First, interpretations of their locational properties, expressed informally by
390 archaeologists, have been statistically expressed and validated. The importance of this should
391 not be understated. Drawing reliable conclusions as to the structure, formation processes, and
392 wider distributional patterns of geometric earthworks is key to building robust models of pre-
393 Columbian social organisation and going beyond the plain-language explanations typically

394 favoured by archaeologists. Additionally, while first-order factors do not explain the structure
395 of the archaeological point pattern on their own, the incorporation of a term to directly address
396 point interaction both provides a model with a good fit for the majority of the variability in the
397 Acre earthworks, and also hints towards resolving an elusive property of these structures.
398 Specifically, this discussion now draws attention to the spatial scale at which patterning among
399 geometric earthworks manifests most strongly (**Fig. 3, bottom right**). What might this imply
400 for the organisation and landscape-level structure of pre-Columbian societies in Acre and
401 western Amazonia more broadly?

402

403 This question is far from trivial, since direct connections have been drawn between the level
404 of regional integration, anthropic landscape history, and the size and elaboration of geometric
405 earthworks (Saunaluoma et al. 2018; Watling et al. 2018). A salient point first underlined by
406 Schaan (2012: 154-155) is that the distribution of the largest earthworks may reflect some
407 aspect of locally marked hierarchies, which interacted at a regional scale through horizontal
408 power relations. The knowledge gained from the point process modelling can be incorporated
409 into further exploratory analyses to this end. In [summary](#): the results demonstrate that
410 significant spatial autocorrelation exists between earthworks in intervals of approximately ~15
411 km, or more precisely, a neighbourhood radius of around 7.5 km (7632 m, to be exact; **Fig.**
412 **3D**).

413

414 Using the neighbourhood radius as a distance threshold (ϵ) in DBSCAN (Density-Based Spatial
415 Clustering of Applications with Noise; see Ester et al. 1996), a data-clustering algorithm
416 particularly suited to point patterns of arbitrary shape and density, groupings of earthworks
417 should emerge that reflect archaeologically meaningful units. Minimum group size is
418 somewhat arbitrarily set at three earthworks, a value that conservatively excludes most sites
419 that are located close to the edges of forest clearances (**Fig. 2A**). This minimum group size
420 threshold approximates a subset of earthworks on the study area boundaries that lack sufficient
421 context for robust interpretations for the time being. Undoubtedly there are more earthworks
422 waiting to be uncovered under the canopy. Finally, potential variation in group membership
423 was also considered by permuting the DBSCAN algorithm through the range of statistically
424 significant clustering (14500 – 15900 m in 100-metre intervals, **Fig. 3D**). This procedure
425 (**Supplementary Fig. 1**) illustrates that varying ϵ primarily alters the relative proportion of
426 “noisy” points with no group. While total number of groups varies between 17 and 21, most
427 shifts in membership occur overwhelmingly among the smallest permissible groups (3

428 earthworks), which either “become noisy” or join larger groups as the distance threshold
 429 decreases or increases, respectively. Most of the groups detected by the algorithm are robust
 430 across thresholds and do not have large alterations to their membership.

431

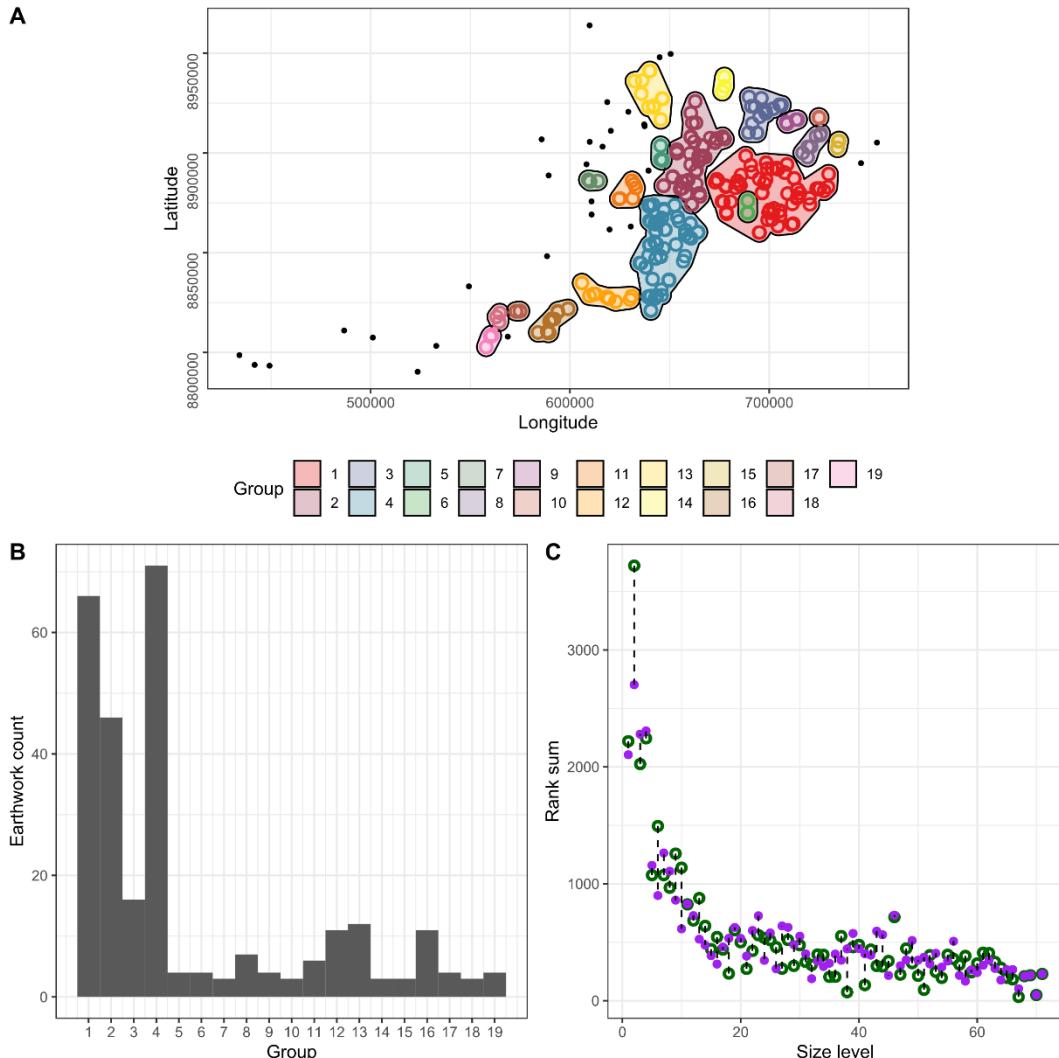


Fig. 4: Non-hierarchical organisation of Amazonian geometric earthworks. **A:** The DBSCAN algorithm defines 19 discrete clusters, highlighted with their convex hulls, with clusters defined at a threshold of $\epsilon = 7632$ metres. Convex hulls do not imply territorial extent, but are presented as a visual aid. Non-clustered points are represented as black dots. **B:** Count of earthworks by group, *minimum group size* = 3. **C:** Differences in rank sums of site area between an ideal hierarchy (green) and the observed distribution (purple) given the same number of earthworks and groups. Further permutation testing indicates that these differences are overwhelmingly likely due to chance rather than within-group hierarchies.

432 The 19 clusters defined among the earthworks at the strongest clustering threshold is shown in
 433 **Fig. 4A**. The groups vary substantially in size from 71 earthworks, enclosing a total of 94.48
 434 ha, to the minimum of three individual sites. The resulting output is visually intuitive, in the
 435 sense that borders between clusters are readily apparent, as well as between the outliers. Groups

436 1, 2, and 4 clearly dominate the clusters in terms of sheer numbers, presenting a relatively
437 skewed distribution (**Fig. 4B**) with a few large groups and many smaller ones. Group 1 wholly
438 surrounds the much smaller Group 6. Nonetheless, the meaningful clusters generated by
439 combining the point process model with DBSCAN, though having highly uneven membership,
440 do not directly inform on whether relative rankings exist within them. Bearing in mind the role
441 of earthwork size in indexing relative site importance (Schaan 2012), if this attribute is a
442 possible marker of hierarchical social relations and community structure, the groupings
443 detected here may provide a useful point of departure for further analyses. To help approximate
444 a solution, the final part of this paper adapts a permutation test developed by Carrero-Pazos et
445 al. (2019) to examine if earthwork area is distributed more unevenly than would be expected
446 by random chance **both between and within groups/clusters**.

447

448 In outline (Carrero-Pazos et al. 2019: 6-7), the method orders all earthworks by size and creates
449 a ranking of site sizes *within* each group: the largest site in each cluster is graded 1, the second
450 largest is graded 2, and so on. The sum of *observed* grades by group is compared to the sum of
451 the *ideal* distribution of grades if a perfect site hierarchy were present, given the same number
452 of ranks. Consistently ideal size distributions within groups are expected to represent internally
453 **ranked social formations**. Here, the number of grades is 71, corresponding to the group size of
454 the largest group. **Fig. 4C** demonstrates that observed rankings overall tend to deviate from the
455 ideal rankings. Notably, size grades 1 and 2 (the largest earthworks, thought to be dominant
456 with each group) lie well below the rank sum that would be expected if a strong hierarchy of
457 sites were present within the groups of earthworks. Below the largest size grades, however, the
458 rank sums tend to exceed the expectations of the ideal ranking, suggesting that small-to-
459 medium earthworks are over-represented relative to the large or exceptionally large earthworks
460 **in the sample**. Initially, at least, an internal hierarchy of monuments is counter-indicated by this
461 test. In other words, if the clusters of **geometric earthworks** can be interpreted as reflecting
462 community organisation, it does not at present seem probable that they were strongly stratified.

463

464 To establish whether the differences between observation and expectation are greater than
465 could occur chance, the size grades can be randomly shuffled independent of group
466 membership. This is performed 999 times. The observed ranking can then be compared to the
467 simulated shuffled rankings, with an associated *p*-value. This procedure indicates that at almost
468 every grade, the distribution of earthwork sizes in the 19 groups (**Fig. 4A**) cannot be
469 distinguished from the null hypothesis of no internal rankings (**Supplementary Table 1**). Two

470 exceptions are at rank 10 and 32, which return a result significant to $p < 0.05$, but no lower.
471 The implication of this specific finding is not straightforward to interpret. The high p -values
472 suggest this is an artefact of the underlying data, particularly in the latter case, as only three
473 groups contain 32 earthworks or more. Overall, it can be suggested that the earthwork groups
474 defined through point pattern analysis are far from hierarchically distributed in terms of their
475 size. While the construction of *individual* earthworks may have been guided by people with
476 the knowledge or authority to do so, there is little evidence at present that centralised control
477 structured the monumental landscape of Acre, either between or within clusters, over time. This
478 has broad-ranging implications for interpreting the formation of this archaeological landscape.
479

480 **5. The structure and organisation of geometric earthworks**

481 Pärssinen et al. (2009) cautioned a decade ago that archaeologists are only just beginning to
482 understand the earthwork builders of the Amazon, and at the time estimated that no more than
483 10% of all sites in the western Amazon had been documented. Far fewer have ever been visited
484 by an archaeologist. Findings such as those presented here are therefore understood to be
485 subject to testing with field data. The approach is deliberately couched in the language of
486 probability, examining problems in Amazonian archaeology in the light of a formalised model-
487 testing approach. This paper has demonstrated that archaeological hypotheses on geometric
488 earthworks are informative and appropriate for direct statistical scrutiny, an increasingly
489 prominent practice in the humanities more broadly (Epstein 2006; Kohler and Van der Leeuw
490 2007) and in landscape archaeology specifically (Bevan et al. 2013). This procedure estimated
491 the relative importance of known variables and detected conditions under which they do not
492 apply, which were controlled for through further testing. Bearing in mind our nascent
493 understanding of western Amazonian societies, the methods provide a rigorous point of
494 departure for extensions and alternative approaches in future work. Social dynamics, whether
495 vertical or horizontal, may be theorised on extensively, but ultimately must be subjected to
496 rigorous scrutiny if they are to improve our understanding of the past (Smith and Schreiber
497 2005: 205).

498
499 In this regard, ‘complexity’ in the abstract is frequently invoked to frame the emergence and
500 endurance of earthwork building as a cultural practice across the Neotropics (Roosevelt 1999;
501 Stahl 2002; Balée and Erickson 2006; Heckenberger and Neves 2009; De Souza et al. 2018),
502 however, operationalising a common definition and testing it in formal, comparative terms has
503 been less forthcoming (Smith and Schreiber 2005). Reflecting on this broader issue of model

504 testing in lowland South America, this paper has argued that archaeological accounts of human-
505 environmental relations would benefit from the systematic application of computational
506 methods. The approach advocated here is not antithetical to frameworks that emphasise the
507 dynamics of social behaviour or political organisation, but build usefully upon existing
508 hypothesis with complementary, probabilistic analytical tools. Point process modelling enables
509 multiple effects to be directly integrated in explanatory frameworks (McMichael et al. 2014a;
510 2014b). Statistically informative gains were achieved through a synthesis of model terms that
511 encapsulated both first- and second-order interaction within the material record. The former
512 class of effects collectively reflected consistent elements of geometric earthwork
513 configurations in relation to the landscape, while the latter stood in as a proxy for potential
514 social processes guiding the placement of these monuments in relation to one another.
515 Achieving a multifactorial fit to data in this manner demonstrates that archaeological models
516 expressed in plain language are tractable to robust statistical analysis. Formal simulation-based
517 models of complex archaeological realities are useful in drawing out salient elements of the
518 material record at the landscape scale (Wobst 1974; Lake 2015), and ought to be standard. This
519 paper provides a point of departure for further advances in point process modelling
520 (**Supplementary Information**), for example of other traditions of earthwork building among
521 the indigenous societies of pre-Columbian South America (Arnold and Prettol 1988;
522 Heckenberger et al. 1999; Roosevelt 1999; Prümers et al. 2006; Walker 2008; Lombardo and
523 Prümers 2010; Prümers 2014; De Souza et al. 2018).

524

525 Although it appears unlikely that time-transgressive or ascribed political status was a factor in
526 behind the construction of geometric earthworks, their distribution and groupings may instead
527 reflect other aspects of society in the pre-Columbian period, as well as biogeographical factors
528 (McMichael et al. 2014a; Watling et al. 2017). Bearing in mind the link that has repeatedly
529 been made between labour mobilisation and earthwork size, it can be tentatively suggested that
530 a demographic dimension to their distribution and degree of elaboration may exist. As noted in
531 the introduction, the period of [geometric](#) earthwork construction spans some two millennia. If
532 a certain degree of population fluctuation is assumed over such a lengthy time span, the
533 interplay between population growth and contraction [would suggest, in theory, that the](#)
534 [presence of larger sites simply reflects](#) there being more people to mobilise at a given point in
535 time, and smaller sites the reverse. On the other hand, Saunaluoma and Virtanen (2015: 36) do
536 not believe there is an association between increased [geometric](#) earthwork elaboration and the
537 advance of time. This link therefore remains somewhat conjectural, and its confirmation or

538 rejection ought to rest on far more extensive sampling of radiocarbon data from both
539 [monumental sites](#) and the few settlement sites that are known in the western Amazon. Future
540 investigations may also benefit from direct estimates of other geometric earthwork parameters,
541 such as the energetics involved in construction: berm height, ditch depth, total volume of soil
542 movement, [and/or](#) new indices of architectural complexity. Such variables, should they be
543 collected, could prove informative in renewed modelling efforts and permutation testing.
544 Additionally, the acquisition of high-resolution chronometric data would enable spatio-
545 temporal point process models to be fitted to archaeological patterns, turning the type of
546 analysis developed here from synchronic to diachronic (Baddeley and Turner 2005).
547 Extensions of the pair-correlation function that test for patterning among categorical subsets of
548 points, usually termed bi- or multi-variate spatial statistics would directly address phenomena
549 such as the reported cline in [geometric](#) earthwork morphology identified by Schaan (2012)
550 along a north-south axis.

551

552 [Returning to the present research: the combined model highlighted hereto-unrecognised spatial](#)
553 [structure among geometric earthworks in Acre. Further examination of these patterns with the](#)
554 [simulation and permutation testing enabled specific inferences on the types of societies that](#)
555 [likely lived in this part of western Amazonia during the Late Holocene. Second, in a broader](#)
556 [methodological and theoretical scope, the analysis provided valuable insight into the](#)
557 [development of spatial analytical approaches in indigenous Amazonian contexts. The basic](#)
558 [null model \(Fig. 3\) illustrates that statistically significant clustering *and* dispersion occurs](#)
559 [between geometric earthworks at a variety of ranges. Furthermore, the additional controls on](#)
560 [the observed pattern \(Fig. 2, Table 1\) produced major insights into where the significant,](#)
561 [“extraordinary” patterns are in the data. Armed with this knowledge, we can affirm that](#)
562 [analysis at arbitrary spatial scales would fail to capture the diversity of the archaeological point](#)
563 [pattern. The results make it clear that fixed-scale summary statistics under a null assumption](#)
564 [of spatial homogeneity will not succeed in adequately describing data that are spatial](#)
565 [palimpsests \(Davies et al. 2016\). With reference to the study of “complexity”, a pervasive](#)
566 [rhetorical device in the archaeology of the Amazon basin, there are clear benefits to using](#)
567 [analytical strategies that can detect and account for non-linear interactions of this nature at](#)
568 [multiple spatial scales.](#)

569

570 On one level, the findings presented here suggest that stratified or ranked societies are unlikely
571 to have been behind the construction of groups of [geometric](#) earthworks in western Amazonia.

572 Patterning within groups, whose definition is derived through robust inference, is more
573 evocative of non-hierarchical societies. It follows that interactions *between* groups of
574 earthwork-builders may have transiently expressed localised authority, materialised through
575 place-making practices in the form of geometric earthworks (Saunaluoma and Virtanen 2015).
576 Emergent forms of political control, in Acre as in many places in the world, were likely
577 historically variable and actively negotiated (Furholt et al. 2019). Ultimately, however, the
578 establishment of such relationships were not necessarily leveraged to form time-transgressive,
579 top-down political control over extended social groups. This accords with theory (Crumley
580 1979, 1995) and the archaeological record of Acre, which has thus far not produced evidence
581 typically associated with assigned social rank, such as structured access to resources, elite
582 burials, or as established here, a clear site hierarchy. Tentatively, an absence of centralised
583 control at any scale may also explain the diversity of [geometric](#) earthwork shapes, degree of
584 elaboration (Saunaluoma et al. 2018), and probably cluster size too. Anthropic changes to
585 vegetation community structure are evident at a landscape level in Acre since at least 4000 cal
586 BP, with major alterations to the relative abundance of comestible and economically useful
587 species (Watling et al. 2017, 2018). [Without falling into unhelpful dichotomies between](#)
588 [intentional niche construction versus accidental forest enrichment, I note that evidence for](#)
589 [systematic, coordinated efforts at environmental modification is also lacking in Acre](#)
590 [\(McMichael et al. 2012; Watling et al. 2017\).](#) This contrasts with neighbouring regions of
591 Amazonia, where both social stratification and large-scale labour coordination are reported in
592 association (Erickson 2006; Lombardo and Prümers 2010; Prümers and Jaimes Betancourt
593 2014). Despite the relative geographic proximity of geometric earthworks to other, diverse
594 earthmoving cultures of the Llanos de Moxos, their different organisational trajectories caution
595 against over-generalising the drivers and mechanisms behind the formation of extensive pre-
596 Columbian built environments without accounting for local trajectories first. In closing, the
597 observed diversity in their configurations and spatial structure likely possesses a chronological
598 element, as well as cultural and possibly demographic dimensions. These issues, highlighted
599 by the fitted point process models presented here, can only be resolved through further
600 excavation and recovery of data on their life histories.

601

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610

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